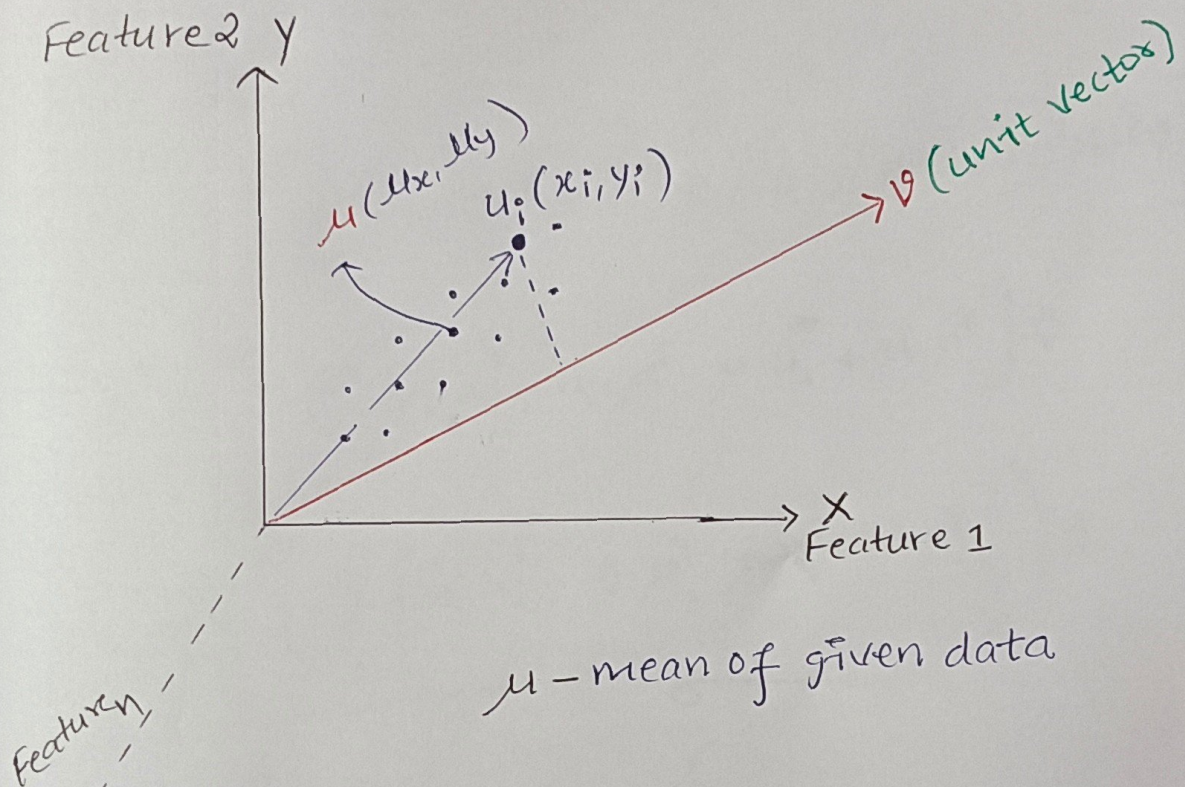


# PCA

Aim:- PCA aims to find directions (unit vectors) along which the projected data exhibits the highest possible variance.

The problem Formulation:-



Now,

$$\text{Variance given data} = \frac{\sum_{i=1}^n (u_i - \mu)^2}{n}$$

Variance of projected data =

$$V = \frac{1}{n} \sum_{i=1}^n (u_i^T \cdot v - \mu^T \cdot v)^2$$

\* Now PCA aims to maximize this variance ( $V$ )

Let's simplify  $V$

$$\begin{aligned} V &= \frac{1}{n} \sum_{i=1}^n (\underbrace{u_i^T \cdot v}_{\downarrow} - \underbrace{\mu^T \cdot v}_{\downarrow}) (\underbrace{u_i^T \cdot v}_{\downarrow} - \underbrace{\mu^T \cdot v}_{\downarrow}) \\ &= \frac{1}{n} \sum_{i=1}^n (v^T \cdot u_i - v^T \cdot \mu) \cdot (u_i^T \cdot v - \mu^T \cdot v) \quad \text{Just interchanged} \\ &= \frac{1}{n} \sum_{i=1}^n \left[ v^T u_i \cdot u_i^T v - v^T u_i \cdot \mu^T v - v^T \mu \cdot u_i^T v + v^T \mu \mu^T v \right] \\ &= \frac{1}{n} \sum_{i=1}^n v^T \left[ u_i \cdot u_i^T - u_i \cdot \mu^T - \mu \cdot u_i^T + \mu \cdot \mu^T \right] v \\ &= v^T \left( \frac{1}{n} \sum_{i=1}^n \left[ u_i \cdot u_i^T - u_i \cdot \mu^T - \mu \cdot u_i^T + \mu \cdot \mu^T \right] \right) v \\ &= v^T C v \end{aligned}$$

Therefore

$V = v^T C v$  → Now our goal is to maximize this variance  $V$  with a constraint  $\|v\| = 1$  ( $\because v$  is a unit vector)

we know  $\|v\| = 1$

$$\begin{aligned}\Rightarrow v^T \cdot v &= |v^T| |v| \cos \theta \\ &= |v^T| |v| \cos(0) \\ &= 1 \cdot 1 = 1\end{aligned}$$

$\therefore$  our constraint is  $v^T \cdot v = 1$

\* Our objective is to maximize variance  $v$  subject to a constraint  $v^T \cdot v = 1$ . Such problems of constrained optimization might be reformulated as unconstrained optimization problems via use of **Lagrangian multipliers**.

**Lagrangian multipliers:-**

If we would like to maximize  $f(x)$  subject to  $g(x) = c$ , we introduce the Lagrange multiplier  $\lambda$  and construct the Lagrangian

$L(x, \lambda)$

$$L(x, \lambda) = f(x) - \lambda(g(x) - c)$$

\* The sign of  $\lambda$  doesn't make any difference.

We then solve the system of equations:

$$\left\{ \begin{array}{l} \frac{\partial L(x, \lambda)}{\partial x} = 0, \\ \frac{\partial L(x, \lambda)}{\partial \lambda} = 0 \end{array} \right\}$$

The Solution:-

In our case, we want to maximize Variance

$v$  subject to a constraint  $v^T \cdot v = 1$ .

$f(x)$

$g(x) = c$

our Lagrangian is:

$$L(v, \lambda) = v^T C v - \lambda (v^T \cdot v - 1)$$

We solve for the stationary points:

$$\frac{\partial L(v, \lambda)}{\partial v} = \frac{\partial}{\partial v} [v^T C v - \lambda (v^T \cdot v - 1)]$$

$$0 = C \frac{\partial}{\partial v} (v^T \cdot v) - \lambda \frac{\partial}{\partial v} (v^T \cdot v)$$

$$\Rightarrow C \frac{\partial}{\partial v} (|v|^2) = \lambda \frac{\partial}{\partial v} (|v|^2) \quad \left[ \begin{array}{l} v^T \cdot v = |v|^2 \\ \therefore |v^T| = |v| \end{array} \right]$$

$$C(x^i) = \lambda(x^i)$$

$$\Rightarrow \boxed{Cv = \lambda v} \quad \text{😊}$$

↳ It is an Eigen Vector Equation

\* The vector we are searching for is an eigen vector, Guys!!

Now,

$$\frac{\partial L(v, \lambda)}{\partial \lambda} = \frac{\partial [v^T C v - \lambda (v^T v - 1)]}{\partial \lambda}$$

$$0 = v^T v - 1 \rightarrow \text{we already know it}$$

let's *Decode* what is  $C$ ?

$$C = \frac{1}{n} \sum_{i=1}^n [u_i \cdot u_i^T - u_i \cdot \mu^T - \mu \cdot u_i^T + \mu \cdot \mu^T]$$

$u_i = \begin{bmatrix} x_i \\ y_i \end{bmatrix}$   $\mu = \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}$  → matrix representation of a vector.

$\mu_x$  - Mean feature 1 (x)

$\mu_y$  = Mean of feature 2 (y)

Now,

$$u_i \cdot u_i^T = \begin{bmatrix} x_i \\ y_i \end{bmatrix} \begin{bmatrix} x_i & y_i \end{bmatrix} = \begin{bmatrix} x_i^2 & x_i y_i \\ y_i x_i & y_i^2 \end{bmatrix}$$

$2 \times 1$                        $1 \times 2$                        $2 \times 2$

$$u_i \cdot \mu^T = \begin{bmatrix} x_i \\ y_i \end{bmatrix} \begin{bmatrix} \mu_x & \mu_y \end{bmatrix} = \begin{bmatrix} x_i \mu_x & x_i \mu_y \\ y_i \mu_x & y_i \mu_y \end{bmatrix}$$

$2 \times 2$

$$\mu \cdot u_i^T = \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix} \begin{bmatrix} x_i & y_i \end{bmatrix} = \begin{bmatrix} \mu_x x_i & \mu_x y_i \\ \mu_y x_i & \mu_y y_i \end{bmatrix}$$

$2 \times 2$

$$\mu \cdot \mu^T = \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix} \begin{bmatrix} \mu_x & \mu_y \end{bmatrix} = \begin{bmatrix} \mu_x^2 & \mu_x \mu_y \\ \mu_y \mu_x & \mu_y^2 \end{bmatrix}$$

$2 \times 2$

$$\therefore C = \frac{1}{n} \sum_{i=1}^n \begin{bmatrix} x_i^2 - x_i \mu_x - \mu_x x_i + \mu_x^2 & x_i y_i - x_i \mu_y - \mu_x y_i + \mu_x \mu_y \\ y_i x_i - y_i \mu_x - \mu_y x_i + \mu_x \mu_y & y_i^2 - 2 \mu_y y_i + \mu_y^2 \end{bmatrix}$$

$$C = \begin{bmatrix} \frac{1}{n} \sum_{i=1}^n (x_i - \mu_x)^2 & \frac{1}{n} \sum_{i=1}^n (x_i - \mu_x)(y_i - \mu_y) \\ \frac{1}{n} \sum_{i=1}^n (x_i - \mu_x)(y_i - \mu_y) & \frac{1}{n} \sum_{i=1}^n (y_i - \mu_y)^2 \end{bmatrix}$$

**Covariance**:- It measures the relationship between two variables.

\* A **positive** covariance indicates that variables tend to change together (both increase or both decrease), while **negative** covariance means they tend to change in opposite directions.

Covariance-matrix

$$[\text{COV}(X, Y)] = \begin{bmatrix} \text{COV}(X, X) & \text{COV}(X, Y) \\ \text{COV}(Y, X) & \text{COV}(Y, Y) \end{bmatrix}$$

matrix

↑  
for 2-D

$$= \begin{bmatrix} \text{Var}(X) & \text{COV}(X, Y) \\ \text{COV}(Y, X) & \text{Var}(Y) \end{bmatrix}$$

we know

$$\text{COV}(X, Y) = \text{COV}(Y, X) = \frac{1}{n} \sum_{i=1}^n (x_i - \mu_x)(y_i - \mu_y)$$

**Covariance matrix** =

$$\begin{bmatrix} \frac{1}{n} \sum_{i=1}^n (x_i - \mu_x)^2 & \frac{1}{n} \sum_{i=1}^n (x_i - \mu_x)(y_i - \mu_y) \\ \frac{1}{n} \sum_{i=1}^n (x_i - \mu_x)(y_i - \mu_y) & \frac{1}{n} \sum_{i=1}^n (y_i - \mu_y)^2 \end{bmatrix}$$

Very Very clearly,

Covariance Matrix = C



I am Very Happy

Guys, I hope you can recollect we use to find the eigen value and eigen vectors of covariance matrix. This is the Reason!!!